EXERCISE 1 (ADA BOOST) Let $(x_i, y_i)_{1 \le i \le n} \in (X \times \{-1, 1\})^n$ be *n* observations and $H = \{h_1, \ldots, h_M\}$ be a set of *M* classifiers, i.e. for all $1 \le i \le M$, $: h_i : X \to \{-1, 1\}$. It is assumed that for each $h \in H$ and that there exist $1 \le i \ne j \le n$ such that $y_i = h(x_i)$ and $y_j \ne h(x_j)$. Let F be the set of all linear combinations of elements of H:

$$\mathsf{F} = \left\{ \sum_{j=1}^{M} \theta_j h_j \, ; \, \theta \in \mathbb{R}^M \right\} \, .$$

Consider the following algorithm. Set $\hat{f}_0 = 0$ and for all $1 \leq m \leq M$,

$$\hat{f}_m = \hat{f}_{m-1} + \beta_m h_{j_m} \quad \text{where} \quad (\beta_m, h_{j_m}) = \operatorname*{argmin}_{h \in \mathsf{H}, \beta \in \mathbb{R}} n^{-1} \sum_{i=1}^n \exp\left\{-y_i \left(\hat{f}_{m-1}(x_i) + \beta h(x_i)\right)\right\} \ .$$

1. Choosing $\omega_i^m = n^{-1} \exp\{-y_i \hat{f}_{m-1}(x_i)\}$, show that

$$n^{-1}\sum_{i=1}^{n}\exp\left\{-y_{i}\left(\hat{f}_{m-1}(x_{i})+\beta h(x_{i})\right)\right\} = \left(e^{\beta}-e^{-\beta}\right)\sum_{i=1}^{n}\omega_{i}^{m}\mathbb{1}_{h(x_{i})\neq y_{i}} + e^{-\beta}\sum_{i=1}^{n}\omega_{i}^{m}.$$

2. For all $1 \leq m \leq M$ and $h \in H$, define

$$\operatorname{err}_{m}(h) = \frac{\sum_{i=1}^{n} \omega_{i}^{m} \mathbb{1}_{h(x_{i}) \neq y_{i}}}{\sum_{i=1}^{n} \omega_{i}^{m}} \,.$$

Prove that

$$h_{j_m} = \mathop{\rm argmin}_{h \in \mathsf{H}} \operatorname{err}_m(h) \quad \text{and} \quad \beta_m = \frac{1}{2} \log \left(\frac{1 - \operatorname{err}_m(h_{j_m})}{\operatorname{err}_m(h_{j_m})} \right) \; .$$

3. Propose an algorithm to compute \hat{f}_M .

EXERCISE 2 (CONSISTENCY OF A SIMPLE RANDOM FOREST) Consider a data set $\mathcal{D}_n = \{(X_i, Y_i) \in [0, 1]^d \times \mathbb{R}, i = 1, ..., n\}$. It is assumed that the (X_i, Y_i) are i.i.d. with the same distribution as (X, Y) where

$$Y = r(X) + \varepsilon,$$

with ε a centered Gaussian noise, independent of X and r a continuous function. Define the following centered random forest estimator:

- 1. Grow M trees as follows:
 - (a) Consider the cell $[0,1]^d$.
 - (b) Select uniformly one variable j^* in $\{1, \ldots, d\}$.
 - (c) Cut the cell at the middle of the j^* -th side, where j^* is the coordinate chosen above.
 - (d) For each of the two resulting cells, repeat (b) (c) if the cell has been cut strictly less than k_n times.
 - (e) For a query point x, the m-th tree outputs the average $\hat{r}_n(x, \Theta_m)$ of the Y_i falling into the same cell as x, where Θ_m is the random variable encoding all selected splitting variables in each cell of the m-th tree.
- 2. For a query point x, the centered forest outputs the average $\hat{r}_{M,n}(x,\Theta_1,\ldots,\Theta_M)$ of the predictions given by the M trees.

Define the infinite random forest estimate $\hat{r}_{\infty,n}$ by considering the random forest estimate defined above and letting $M \to \infty$, that is

$$\hat{r}_{\infty,n}(x) = \mathbb{E}_{\Theta}[\hat{r}_n(x,\Theta)],$$

where \mathbb{E}_{Θ} is the expectation with respect to Θ only. For a tree built with the randomness Θ , we let $A_n(x, \Theta)$ be the cell containing x and $N_n(x, \Theta)$ be the number of observations falling into $A_n(x, \Theta)$. We want to prove the following theorem:

- **Theorem 1.** Assume that $k_n \to \infty$ is such that $2^{k_n}/n \to 0$, as $n \to \infty$. Then the random forest fulfills $\mathbb{E}[(\hat{r}_{\infty,n}(X) r(X))^2] \to 0$, where X is independent of $(X_i, Y_i)_{i=1,...,n}$ with the same distribution as the X_i on $[0, 1]^d$.
 - 1. Prove that there exists weights $W_{ni}(x,\Theta)$ and $W_{ni}^{\infty}(x)$, $1 \leq i \leq n$, such that

$$\hat{r}_n(x,\Theta) = \sum_{i=1}^n W_{ni}(x,\Theta)Y_i, \quad \text{and} \quad \hat{r}_{\infty,n}(x) = \sum_{i=1}^n W_{ni}^\infty(x)Y_i.$$

In this context, Stone's Theorem states that the random tree estimate $\hat{r}_n(x,\Theta)$ fulfills

$$\lim_{n \to \infty} \mathbb{E}\left[(\hat{r}_n(X, \Theta) - r(X))^2 \right] = 0,$$

as soon as the two following conditions are satisfied

(i) $\mathbb{E}[\operatorname{diam}(A_n(X,\Theta))] \to 0$, as $n \to \infty$, where the diameter of any cell A is defined as

$$\operatorname{diam}(A) = \sup_{x,z \in A} \|x - z\|_2.$$

- (*ii*) $N_n(X, \Theta) \to \infty$ in probability, as $n \to \infty$.
 - 2. Let $x \in [0,1]^d$. What is the distribution of the number of cuts along the coordinate $j \in \{1, \ldots, d\}$ in the cell $A_n(x, \Theta)$?
 - 3. Check that, for all $x \in [0,1]^d$ and $j \in \{1,\ldots,d\}$,

$$\mathbb{E}\left[\sup_{z\in A_n(x,\Theta)} z_j - \inf_{z\in A_n(x,\Theta)} z_j\right] = \left(1 - \frac{1}{2d}\right)^{k_n}$$

- 4. Prove that (i) holds for a random centered tree.
- 5. We denote by $A_1, \ldots, A_{2^{k_n}}$ the 2^{k_n} cells and by N_ℓ the number of points among X, X_1, \ldots, X_n which fall into A_ℓ . Show that for $\ell \in \{1, \ldots, 2^{k_n}\}$,

$$\mathbb{P}\left(X \in A_{\ell} | N_{\ell}, \Theta\right) = \frac{N_{\ell}}{n+1}.$$

Conclude that for every integer t > 0,

$$\mathbb{P}\left(N_n(X,\Theta) \leqslant t\right) \leqslant t 2^{k_n} / (n+1)$$

and hence that (ii) holds for a random centered tree.

- 6. Prove that the infinite centered random forest fulfills $\mathbb{E}[(\hat{r}_{\infty,n}(X) r(X))^2] \to 0$, as $n \to \infty$.
- 7. Assume that the noise ε is Gaussian with variance $\sigma^2 > 0$. Thus,

$$\mathbb{E}\left[\max_{1\leqslant i\leqslant n}\varepsilon_i^2\right]\leqslant \sigma^2(1+4\log n).$$

Find a condition on the number M_n of trees such that the finite centered random forest fulfills

$$\lim_{n \to \infty} \mathbb{E}[(\hat{r}_{M_n,n}(X,\Theta_1,\ldots,\Theta_{M_n}) - r(X))^2] = 0.$$