

Simple Linear Regression: exercises

Exercise 1 (Estimation) *By minimising*

$$S(\beta_1, \beta_2) = \sum_{i=1}^n (y_i - \beta_1 - \beta_2 x_i)^2.$$

Find the expression of $\hat{\beta}_1$ and $\hat{\beta}_2$.

Calculate the derivatives according to β_1 and β_2 , we get

$$\begin{cases} \frac{\partial S(\hat{\beta}_1, \hat{\beta}_2)}{\partial \beta_1} = -2 \sum_{i=1}^n (y_i - \hat{\beta}_1 - \hat{\beta}_2 x_i) = 0, \\ \frac{\partial S(\hat{\beta}_1, \hat{\beta}_2)}{\partial \beta_2} = -2 \sum_{i=1}^n x_i (y_i - \hat{\beta}_1 - \hat{\beta}_2 x_i) = 0. \end{cases}$$

The first equation yields

$$\hat{\beta}_1 n + \hat{\beta}_2 \sum_{i=1}^n x_i = \sum_{i=1}^n y_i$$

and we have the estimator for the intercept

$$\hat{\beta}_1 = \bar{y} - \hat{\beta}_2 \bar{x}, \tag{1}$$

where $\bar{x} = \sum_{i=1}^n x_i/n$. The second equation yields

$$\hat{\beta}_1 \sum_{i=1}^n x_i + \hat{\beta}_2 \sum_{i=1}^n x_i^2 = \sum_{i=1}^n x_i y_i.$$

Replacing $\hat{\beta}_1$ by its expression (1) we can rewrite it as

$$\hat{\beta}_2 = \frac{\sum x_i y_i - \sum x_i \bar{y}}{\sum x_i^2 - \sum x_i \bar{x}},$$

Making use of the fact that the sum $\sum (x_i - \bar{x})$ is zero, we can rewrite the expression for the gradient of the straight line as

$$\hat{\beta}_2 = \frac{\sum x_i (y_i - \bar{y})}{\sum x_i (x_i - \bar{x})} = \frac{\sum (x_i - \bar{x})(y_i - \bar{y})}{\sum (x_i - \bar{x})(x_i - \bar{x})} = \frac{\sum (x_i - \bar{x}) y_i}{\sum (x_i - \bar{x})^2}. \tag{2}$$

To obtain this result, we assume that there are at least two points with different x axis coordinates.

Exercise 2 (Bias) *Considering the model*

$$Y_i = \beta_1 + \beta_2 x_i + \varepsilon_i$$

and under the assumption $\mathcal{H}_2 : \mathbb{E}(\varepsilon_i) = 0$, for $i = 1, \dots, n$ and $\text{Cov}(\varepsilon_i, \varepsilon_j) = \delta_{ij}\sigma^2$, evaluate the bias of $\hat{\beta}_2$ and $\hat{\beta}_1$.

Replace y_i by $\beta_1 + \beta_2 x_i + \varepsilon_i$ and get

$$\begin{aligned} \hat{\beta}_2 &= \frac{\sum (x_i - \bar{x}) y_i}{\sum (x_i - \bar{x})^2} = \frac{\beta_1 \sum (x_i - \bar{x}) + \beta_2 \sum x_i (x_i - \bar{x}) + \sum (x_i - \bar{x}) \varepsilon_i}{\sum (x_i - \bar{x})^2} \\ &= \beta_2 + \frac{\sum (x_i - \bar{x}) \varepsilon_i}{\sum (x_i - \bar{x})^2}. \end{aligned}$$

The last term is 0 by construction.

$$\text{Now } \mathbb{E}(\hat{\beta}_1) = \mathbb{E}(\bar{y}) - \bar{x} \mathbb{E}(\hat{\beta}_2) = \beta_1 + \bar{x} \beta_2 - \bar{x} \beta_2 = \beta_1.$$

Exercise 3 (Variance) *Considering the model*

$$Y_i = \beta_1 + \beta_2 x_i + \varepsilon_i$$

and under the assumption $\mathcal{H}_2 : \mathbb{E}(\varepsilon_i) = 0$, for $i = 1, \dots, n$ and $\text{Cov}(\varepsilon_i, \varepsilon_j) = \delta_{ij}\sigma^2$, evaluate the variance of $\hat{\beta}_2$ and $\hat{\beta}_1$.

Start with $V(\hat{\beta}_2)$ since the only random variable is ε we have

$$\begin{aligned} V(\hat{\beta}_2) &= V\left(\beta_2 + \frac{\sum (x_i - \bar{x}) \varepsilon_i}{\sum (x_i - \bar{x})^2}\right) \\ &= V\left(\frac{\sum (x_i - \bar{x}) \varepsilon_i}{\sum (x_i - \bar{x})^2}\right) = \frac{V(\sum (x_i - \bar{x}) \varepsilon_i)}{[\sum (x_i - \bar{x})^2]^2} \\ &= \frac{\sum_{i,j} (x_i - \bar{x})(x_j - \bar{x}) \text{Cov}(\varepsilon_i, \varepsilon_j)}{[\sum (x_i - \bar{x})^2]^2}. \end{aligned}$$

Since $\text{Cov}(\varepsilon_i, \varepsilon_j) = \delta_{ij}\sigma^2$ we have

$$V(\hat{\beta}_2) = \frac{\sum_i (x_i - \bar{x})^2 \sigma^2}{[\sum_i (x_i - \bar{x})^2]^2} = \frac{\sigma^2}{\sum (x_i - \bar{x})^2}.$$

Now

$$\begin{aligned} V(\hat{\beta}_1) &= V(\bar{y} - \hat{\beta}_2 \bar{x}) = V(\bar{y}) + V(\bar{x} \hat{\beta}_2) - 2 \text{Cov}(\bar{y}, \hat{\beta}_2 \bar{x}) \\ &= V\left(\frac{\sum y_i}{n}\right) + \bar{x}^2 \frac{\sigma^2}{\sum (x_i - \bar{x})^2} - 2 \bar{x} \text{Cov}(\bar{y}, \hat{\beta}_2) \\ &= \frac{\sigma^2}{n} + \bar{x}^2 \frac{\sigma^2}{\sum (x_i - \bar{x})^2} - 2 \bar{x} \sum_i \text{Cov}(\bar{y}, \hat{\beta}_2). \end{aligned}$$

Since

$$\begin{aligned}
\text{Cov}(\bar{y}, \hat{\beta}_2) &= \frac{1}{n} \text{Cov} \left(\sum_i (\beta_1 + \beta_2 x_i + \varepsilon_i), \frac{\sum_j (x_j - \bar{x}) \varepsilon_j}{\sum_j (x_j - \bar{x})^2} \right) \\
&= \frac{1}{n} \sum_i \text{Cov} \left(\varepsilon_i, \frac{\sum_j (x_j - \bar{x}) \varepsilon_j}{\sum_j (x_j - \bar{x})^2} \right) \\
&= \frac{1}{\sum_j (x_j - \bar{x})^2} \sum_i \frac{1}{n} \text{Cov} \left(\varepsilon_i, \sum_j (x_j - \bar{x}) \varepsilon_j \right) \\
&= \frac{\sigma^2 \frac{1}{n} \sum_i (x_i - \bar{x})}{\sum_j (x_j - \bar{x})^2} = 0.
\end{aligned}$$

We get

$$V(\hat{\beta}_1) = \frac{\sigma^2}{n} + \bar{x}^2 \frac{\sigma^2}{\sum (x_i - \bar{x})^2} = \frac{\sigma^2 \sum x_i^2}{n \sum (x_i - \bar{x})^2}.$$

Exercice 4 (Covariance) *Considering the model*

$$Y_i = \beta_1 + \beta_2 x_i + \varepsilon_i$$

and under the assumption $\mathcal{H}_2 : \mathbb{E}(\varepsilon_i) = 0$, for $i = 1, \dots, n$ and $\text{Cov}(\varepsilon_i, \varepsilon_j) = \delta_{ij} \sigma^2$, evaluate the covariance of $\hat{\beta}_2$ and $\hat{\beta}_1$.

We have

$$\text{Cov}(\hat{\beta}_1, \hat{\beta}_2) = \text{Cov}(\bar{y} - \hat{\beta}_2 \bar{x}, \hat{\beta}_2) = \text{Cov}(\bar{y}, \hat{\beta}_2) - \bar{x} V(\hat{\beta}_2) = -\frac{\sigma^2 \bar{x}}{\sum (x_i - \bar{x})^2}.$$

Exercice 5 (Sum of residuals) *Show that in a simple linear regression model the sum of the residuals is zero.*

Just write

$$\sum_i \hat{\varepsilon}_i = \sum_i (y_i - \bar{y} + \hat{\beta}_2 \bar{x} - \hat{\beta}_2 x_i) = \sum_i (y_i - \bar{y}) - \hat{\beta}_2 \sum_i (x_i - \bar{x}) = 0.$$