

EXERCISE 1 (REFRESHER ON MATRICES)

1. Let \mathbf{A} be a $n \times d$ matrix with real entries. Show that $\text{range}(\mathbf{A}) = \text{range}(\mathbf{A}\mathbf{A}^T)$.

$$\text{or } \underbrace{\mathbf{A}}_{\in \mathbb{R}^{n \times d}} \quad ? \quad \text{range}(\mathbf{A}) : \text{range}(\mathbf{A}\mathbf{A}^T).$$

$$\text{Ker}(\mathbf{A}\mathbf{A}^T) \supseteq \text{Ker}(\mathbf{A}^T).$$

$$\begin{aligned} & \text{if } x \in \text{Ker}(\mathbf{A}^T), \text{ then: } \underbrace{\mathbf{A}^T x}_{=0} = 0 \quad \text{then } x \in \text{Ker}(\mathbf{A}\mathbf{A}^T). \\ & \text{if } x \in \text{Ker}(\mathbf{A}\mathbf{A}^T), \text{ then: } \underbrace{x^T \mathbf{A}^T x}_{=0} = 0 = \|[\mathbf{A}^T x]\|^2 \\ & \Rightarrow \mathbf{A}^T x = 0 \Rightarrow x \in \text{Ker}(\mathbf{A}^T). \end{aligned}$$

$$\text{Dm: } \text{Ker}(\mathbf{A}\mathbf{A}^T) = \text{Ker}(\mathbf{A}^T).$$

$$\text{Eng: } [\text{Ker}(\mathbf{A}^T)]^\perp = \text{Range}(\mathbf{A}). \quad \text{Effekt: } \text{Range}(\mathbf{A})^\perp = \text{Ker}(\mathbf{A}^T).$$

$$\begin{aligned} & \mathbf{A} = [A_1, \dots, A_d] \\ & x \in \text{Range}(\mathbf{A})^\perp \Leftrightarrow \left[\begin{array}{c} \mathbf{A}^T x \\ A_1^T x \\ \vdots \\ A_d^T x \end{array} \right] = 0 = \mathbf{A}^T x \\ & \Leftrightarrow x \in \text{Ker}(\mathbf{A}^T). \end{aligned}$$

$$\text{Ans: } \text{Range}(\mathbf{A}) = \left\{ \text{Ker}(\mathbf{A}^T) \right\}^\perp = \left\{ \text{Ker}(\mathbf{A}\mathbf{A}^T) \right\}^\perp = \text{Range}(\mathbf{A}\mathbf{A}^T).$$

2. Let $\{U_k\}_{1 \leq k \leq r}$ be a family of r orthonormal vectors of \mathbb{R}^n . Show that $\sum_{k=1}^r U_k U_k^T$ is the matrix associated with the orthogonal projection onto $\mathbf{H} = \{\sum_{k=1}^r \alpha_k U_k ; \alpha_1, \dots, \alpha_r \in \mathbb{R}\}$. Deduce that if \mathbf{A} is a $n \times d$ matrix with real entries such that each column of \mathbf{A} is in \mathbf{H} , then,

$$\left(\sum_{k=1}^r U_k U_k^T \right) \mathbf{A} = \mathbf{A}.$$

$$P_{\mathbf{H}}^{\perp}(x) = \sum_{k=1}^r \underbrace{\langle x, U_k \rangle}_{U_k^T x} \underbrace{U_k}_{M_k} = \underbrace{\left(\sum_{k=1}^r U_k U_k^T \right)}_{\mathbf{A}} x.$$

$$\mathbf{A} = [A_1, \dots, A_d].$$

$$\left(\sum_{k=1}^d U_k U_k^T \right) \left([A_1, \dots, A_d] \right) = \left(\underbrace{\sum_{k=1}^d U_k U_k^T A_1}_{A_1}, \dots, \underbrace{\sum_{k=1}^d U_k U_k^T A_d}_{A_d} \right)$$

$$\boxed{\sum_{k=1}^d U_k U_k^T A = \mathbf{A}}.$$

3. Let $p < d$ and $\mathbf{B} \in \mathbb{R}^{d \times p}$ such that $\mathbf{B}^\top \mathbf{B} = I_p$. Let us denote $\mathbf{B} = (b_{ij})_{\substack{1 \leq i \leq d \\ 1 \leq j \leq p}}$ the components of \mathbf{B} and for all $i \in [1, d]$, $\alpha_i = \sum_{j=1}^p b_{ij}^2$. Show that $\sum_{i=1}^d \alpha_i = p$ and $\alpha_i \leq 1$.

$$\mathbf{B} \in M_{d,p} \Leftrightarrow \mathbf{B} \in \mathbb{R}^{d \times p}. \quad \mathbf{B} = [\mathbf{B}_1, \dots, \mathbf{B}_p] \in \mathbb{R}^{d \times p}, \quad \mathbf{B}^\top \mathbf{B} = I_p.$$

$$\tilde{\mathbf{B}} = [\mathbf{B}_{p+1}, \dots, \mathbf{B}_d]$$

$\tilde{\mathbf{B}} = [\mathbf{B}_1, \dots, \mathbf{B}_p]$

↳ $[\mathbf{B}, \tilde{\mathbf{B}}] \in \mathbb{R}^{d \times d}$ und \mathbb{R}^d

$$\sum_{i=1}^d \alpha_i = \sum_{i=1}^d \sum_{j=1}^p b_{ij}^2 = \text{tr}(\mathbf{B}^\top \mathbf{B}) = p.$$

$$= \sum_{i=1}^d \left(\underbrace{\sum_{j=1}^p b_{ij}^2}_{\| \mathbf{B}_i \|_F^2} \right) = p.$$

$$\mathbf{B} = (\mathbf{B}_1, \dots, \mathbf{B}_p)$$

$$\alpha_i = \sum_{j=1}^p b_{ij}^2 \leq \sum_{j=1}^d b_{ij}^2 = 1.$$

$$\underline{\mathbf{B}} \underline{\mathbf{B}}^\top = \underline{I} \Leftrightarrow \underline{\mathbf{B}}^\top = \underline{\mathbf{B}}^{-1} \Leftrightarrow \underline{\mathbf{B}}^\top \underline{\mathbf{B}} = \underline{I}.$$

EXERCISE 2 (PRINCIPAL COMPONENT ANALYSIS) Principal component analysis is a multivariate technique which aims at analyzing the statistical structure of high dimensional dependent observations by representing data using orthogonal variables called *principal components*. Reducing the dimensionality of the data is motivated by several practical reasons such as improving computational complexity. Let $(X_i)_{1 \leq i \leq n}$ be i.i.d. random variables in \mathbb{R}^d and consider the matrix $\mathbf{X} \in \mathbb{R}^{n \times d}$ such that the i -th row of \mathbf{X} is the observation X_i^\top . In this exercise, it is assumed that data are preprocessed so that the columns of \mathbf{X} are centered. This means that for all $1 \leq k \leq d$, $\sum_{i=1}^n X_{i,k} = 0$. Let Σ_n be the empirical covariance matrix:

$$\Sigma_n = n^{-1} \sum_{i=1}^n X_i X_i^\top.$$

Principal Component Analysis aims at reducing the dimensionality of the observations $(X_i)_{1 \leq i \leq n}$ using a *compression* matrix $\mathbf{U} \in \mathbb{R}^{d \times p}$ with orthonormal columns with $p \leq d$ so that for each $1 \leq i \leq n$, $\mathbf{U}^\top X_i$ is a low dimensional representation of X_i . The original observation may then be partially recovered using \mathbf{U} . Principal Component Analysis computes \mathbf{U} using the least squares approach:

$$\mathbf{U}_* \in \underset{\substack{\mathbf{U} \in \mathbb{R}^{d \times p} \\ \mathbf{U}^\top \mathbf{U} = I_p}}{\operatorname{argmin}} \sum_{i=1}^n \|X_i - \mathbf{U} \mathbf{U}^\top X_i\|^2, \quad \mathbf{U} := [u_1, \dots, u_p] \quad \mathbf{H} := \operatorname{Range}(\mathbf{U})!$$

1. Prove that for all matrix $\mathbf{A} \in \mathbb{R}^{n \times d}$ with rank r , there exist $\sigma_1 \geq \dots \geq \sigma_r > 0$ such that

$$\mathbf{A} = \sum_{k=1}^r \sigma_k u_k v_k^\top, \quad \begin{matrix} \mathbf{u}_k \in \mathbb{R}^n \\ \mathbf{v}_k \in \mathbb{R}^d \end{matrix}$$

where $\{u_1, \dots, u_r\} \subset \mathbb{R}^n$ and $\{v_1, \dots, v_r\} \subset \mathbb{R}^d$ are two families of orthonormal vectors. The vectors $\{u_1, \dots, u_r\}$ (resp. $\{v_1, \dots, v_r\}$) are the left-singular (resp. right-singular) vectors associated with $\{\sigma_1, \dots, \sigma_r\}$, the singular values of \mathbf{A} . If \mathbf{U} denotes the $\mathbb{R}^{n \times r}$ matrix with columns given by $\{u_1, \dots, u_r\}$ and \mathbf{V} denotes the $\mathbb{R}^{d \times r}$ matrix with columns given by $\{v_1, \dots, v_r\}$, then the singular value decomposition of \mathbf{A} may also be written as

$$\mathbf{A} = \mathbf{U} \mathbf{D}_r \mathbf{V}^\top,$$

where $\mathbf{D}_r = \operatorname{diag}(\sigma_1, \dots, \sigma_r)$. Then, $\mathbf{A}^\top \mathbf{A}$ and $\mathbf{A} \mathbf{A}^\top$ are positive semidefinite such that

$$\mathbf{A}^\top \mathbf{A} = \mathbf{V} \mathbf{D}_r^2 \mathbf{V}^\top \quad \text{and} \quad \mathbf{A} \mathbf{A}^\top = \mathbf{U} \mathbf{D}_r^2 \mathbf{U}^\top.$$

In the framework of this exercise, $n \Sigma_n = \mathbf{X}^\top \mathbf{X}$ so that diagonalizing $n \Sigma_n$ is equivalent to computing the singular value decomposition of \mathbf{X} .

$$\begin{aligned} \mathbf{X}^\top \mathbf{X} &= \sum_{h=1}^n \sigma_h u_h v_h^\top = \underbrace{[u_1, \dots, u_n]}_{\mathbf{U}} \underbrace{\begin{pmatrix} \sigma_1 & & 0 \\ 0 & \ddots & \sigma_n \end{pmatrix}}_{\mathbf{D}_r} \underbrace{[v_1^\top, \dots, v_n^\top]}_{\mathbf{V}^\top} \\ &= [\sigma_1 u_1, \sigma_2 u_2, \dots, \sigma_n u_n] \begin{pmatrix} v_1^\top \\ \vdots \\ v_n^\top \end{pmatrix} = \sum_{i=1}^n \sigma_i u_i v_i^\top \end{aligned}$$

Soit $\mathbf{A} \in M_{m,d}(\mathbb{R})$, $\mathbf{A} \mathbf{A}^\top$ est symétrique, n.s. \oplus ($\mathbf{u}^\top \mathbf{A} \mathbf{A}^\top \mathbf{u} = \|\mathbf{A} \mathbf{u}\|^2 \geq 0$)

$$\mathbf{A}^\top \mathbf{A} = \mathbf{U} \begin{pmatrix} \lambda_1 & & 0 \\ 0 & \ddots & \lambda_m \end{pmatrix} \mathbf{U}^\top = \sum_{i=1}^m \lambda_i u_i u_i^\top \quad \text{ai : } \mathbf{U} := (u_1, \dots, u_n). \quad (\lambda_i)_{1 \leq i \leq m} \text{ Bon de vect propres de } \mathbf{A} \mathbf{A}^\top.$$

or : $\operatorname{rang}(\mathbf{A}) = n = \operatorname{rang}(\mathbf{A} \mathbf{A}^\top)$. donc : $\lambda_1 > \dots > \lambda_r > \lambda_{r+1} = \dots = \lambda_n$.

$$\mathbf{A} \mathbf{A}^\top = \sum_{i=1}^n \lambda_i u_i u_i^\top, \quad \mathbf{A} \mathbf{A}^\top \mathbf{u}_i = \lambda_i \mathbf{u}_i \quad \forall i \in \{1, \dots, n\}.$$

$$A^T A \underbrace{A^T u_i}_{\lambda_i} = \lambda_i \underbrace{A^T u_i}$$

$$v_i = \gamma_i A^T u_i$$

$$\text{En prenant } v_i = \frac{A^T u_i}{\sqrt{\lambda_i}} -$$

$$v_i^T v_i = \gamma_i^2 u_i^T \underbrace{A A^T u_i}_{\lambda_i u_i} = \gamma_i^2 \lambda_i = 1 \quad \gamma_i = 1/\sqrt{\lambda_i}$$

On a : $\boxed{A^T A v_i = \lambda_i v_i}$, v_i BON de vect. propres associé à λ_i .
avec $A^T A$.

$$\text{Polarisation: } \boxed{\sigma_i = \sqrt{\lambda_i}}$$

$$\sum_{i=1}^n \sigma_i u_i v_i^T = \sum_{i=1}^n \sqrt{\lambda_i} u_i \left(\frac{A^T u_i}{\sqrt{\lambda_i}} \right)^T = \underbrace{\left(\sum_{i=1}^n u_i u_i^T \right)}_{P \perp \text{Range}(A^T)} A = \boxed{A}.$$

$$\underbrace{\left(\sum_{i=1}^n u_i u_i^T \right) A}_{P \perp \text{Im}(A)} = A.$$

$\text{Range}(A)$.

2. Prove that solving the PCA least squares optimization problem boils down to computing

$$\mathbf{U}_* \in \underset{\mathbf{U} \in \mathbb{R}^{d \times p}, \mathbf{U}^T \mathbf{U} = \mathbf{I}_p}{\operatorname{argmax}} \{\operatorname{Trace}(\mathbf{U}^T \Sigma_n \mathbf{U})\}.$$

$$\mathbf{U}_* \in \underset{\substack{\mathbf{U} \in \mathbb{R}^{d \times p} \\ \mathbf{U}^T \mathbf{U} = \mathbf{I}_p}}{\operatorname{argmin}} \sum_{i=1}^m \underbrace{\| \mathbf{x}_i - \mathbf{U} \mathbf{U}^T \mathbf{x}_i \|_2^2}_{\| \mathbf{x}_i \|_2^2 - 2 \underbrace{\langle \mathbf{x}_i, \mathbf{U} \mathbf{U}^T \mathbf{x}_i \rangle}_{\mathbf{x}_i^T \mathbf{U} \mathbf{U}^T \mathbf{x}_i} + \underbrace{\mathbf{x}_i^T \mathbf{U} \mathbf{U}^T \mathbf{U} \mathbf{U}^T \mathbf{x}_i}_{\mathbf{x}_i^T \mathbf{U} \mathbf{U}^T \mathbf{x}_i}}.$$

$$= \underset{\substack{\mathbf{U}^T \mathbf{U} = \mathbf{I}_p \\ \mathbf{U} \in \mathbb{R}^{d \times p}}}{\operatorname{argmax}} \sum_{i=1}^m \underbrace{\mathbf{x}_i^T \mathbf{U} \mathbf{U}^T \mathbf{x}_i}_{= \operatorname{tr}(\mathbf{x}_i^T \mathbf{U} \mathbf{U}^T \mathbf{x}_i)}.$$

$$= \operatorname{tr}(\mathbf{U}^T \mathbf{x}_i \mathbf{x}_i^T \mathbf{U}).$$

$$= \operatorname{argmax}_{\substack{\mathbf{U} \in \mathbb{R}^{d \times p} \\ \mathbf{U}^T \mathbf{U} = \mathbf{I}_p}} \operatorname{tr}(\mathbf{U}^T (\sum_{i=1}^m \mathbf{x}_i \mathbf{x}_i^T) \mathbf{U}).$$

$$\operatorname{argmax}_{\mathbf{U} \in \mathbb{R}^{d \times p}} \operatorname{tr}(\mathbf{U}^T \Sigma_n \mathbf{U}).$$

$$\mathbf{U}^T \mathbf{U} = \mathbf{I}_p.$$

3. Let $\lambda_1 \geq \dots \geq \lambda_d$ be real numbers and denote $f: \alpha \in \mathbb{R}^d \mapsto \sum_{i=1}^d \alpha_i \lambda_i$. Show that

$$\sup \left\{ f(\alpha) : \alpha \in [0, 1]^d, \sum_{i=1}^d \alpha_i = p \right\}$$

is attained for $\alpha^* = (\mathbf{1}_{i \leq p})_{1 \leq i \leq d}$.

4. Let $\{\vartheta_1, \dots, \vartheta_d\}$ be orthonormal eigenvectors associated with the eigenvalues $\lambda_1 \geq \dots \geq \lambda_d$ of Σ_n . Prove that a solution to the PCA least squares optimization problem is given by the matrix \mathbf{U}_* with columns $\{\vartheta_1, \dots, \vartheta_p\}$.

3). $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_d$.

$$\sum_{i=1}^p \alpha_i \lambda_i \leq \sum_{i=1}^p \alpha_i \lambda_i + \lambda_p \left(\underbrace{\sum_{i=p+1}^d \alpha_i}_{\sum_{i=1}^p \alpha_i} \right) = \sum_{i=1}^p \left(\alpha_i \lambda_i + \underbrace{\lambda_p (1 - \alpha_i)}_{\leq \lambda_i} \right) \leq \sum_{i=1}^p \lambda_i$$

4) $\pi_q: \frac{u^*}{\parallel u \parallel} \in \operatorname{Argmax}_{\substack{u \in \mathbb{R}^{d \times p} \\ (v_1, \dots, v_p) \\ u^T u = I_p}} \text{tr}(u^T \Sigma_m u)$.

qui: v_i : vecteur propre de Σ_m associé à λ_i (avec $\lambda_1 \geq \dots \geq \lambda_d$).

$$\Sigma_m = \underbrace{V D_m V^T}_{= \sum_{i=1}^d \lambda_i V_i V_i^T} \quad \text{où: } \begin{cases} D_m: \operatorname{diag}(\lambda_1, \dots, \lambda_d), \\ V = (V_1, \dots, V_d) \end{cases} \quad \text{où: } (V_i \mid 1 \leq i \leq d) \text{ : BON de vecteurs propres de } \Sigma_m.$$

$$\begin{aligned} \text{tr}(u^T \Sigma_m u) &= \text{tr}\left(\underbrace{u^T V}_{\in \mathbb{R}^{p \times p}} \underbrace{D_m V^T}_{B^T D_m B} \bar{u}\right) = \sum_{i=1}^p \sum_{k=1}^d \underbrace{(B^T)_{ik}}_{B_{ki}} \lambda_k \cdot B_{ki} \\ &= \sum_{i=1}^p \underbrace{\sum_{k=1}^d B_{ki}^2}_{\lambda_{ik}} \lambda_k. \quad \underbrace{B^T B}_{= u^T V \frac{V^T}{I_d} u} : u^T u = I_p \\ &= \sum_{k=1}^d \underbrace{\sum_{i=1}^p B_{ki}^2}_{\lambda_{ik}} \lambda_k. \end{aligned}$$

Pour Ex 1, Q3. $\sum_{i=1}^p \alpha_i = p$, $\alpha_i \in [0, 1]$

Pour Ex 2, Q3. $\sum_{k=1}^d \alpha_k \lambda_k \leq \underbrace{\sum_{k=1}^p \lambda_k}_{= \sum_{h=1}^d \alpha_h' \lambda_h} = \sum_{h=1}^d \alpha_h' \lambda_h$ où $\alpha_h' = \begin{cases} 1 & h \in [1, p] \\ 0 & \text{sinon} \end{cases}$

$$\begin{aligned} \text{En prenant } U_* &= (V_2, \dots, V_p). \quad V = (U_*, \tilde{V}) \quad \text{où } \tilde{V} = (V_{p+1}, \dots, V_d). \\ \text{tr}(U_*^T V D_m V^T U_*) &= \text{tr}\left(\underbrace{U_*^T (U_*, \tilde{V})}_{= (I_p, 0)} D_m \begin{pmatrix} U_*^T \\ \tilde{V} \end{pmatrix} U_*\right) \\ &= \text{tr}\left((I_p, 0) D_m \begin{pmatrix} I_p \\ 0 \end{pmatrix}\right) = \sum_{i=1}^p \lambda_i \end{aligned}$$

5. For any dimension $1 \leq p \leq d$, let \mathcal{F}_d^p be the set of all vector subspaces of \mathbb{R}^d with dimension p . Consider the linear span V_p defined as

$$V_p \in \operatorname{argmin}_{V \in \mathcal{F}_d^p} \sum_{i=1}^n \|X_i - \pi_V(X_i)\|^2,$$

where π_V is the orthogonal projection onto the linear span V . Prove that $V_1 = \operatorname{span}\{v_1\}$ where

$$v_1 \in \operatorname{argmax}_{v \in \mathbb{R}^d; \|v\|=1} \sum_{i=1}^n \langle X_i, v \rangle^2.$$

6. For all $2 \leq p \leq d$, following the same steps, prove that a solution to the optimization problem is given by $V_p = \operatorname{span}\{v_1, \dots, v_p\}$ where

$$\underbrace{v_1 \in \operatorname{argmax}_{v \in \mathbb{R}^d; \|v\|=1} \sum_{i=1}^n \langle X_i, v \rangle^2}_{\text{and for all } 2 \leq k \leq p, \quad v_k \in \operatorname{argmax}_{\substack{v \in \mathbb{R}^d; \|v\|=1; \\ v \perp v_1, \dots, v \perp v_{k-1}}} \sum_{i=1}^n \langle X_i, v \rangle^2}. \quad (1)$$

7. Prove that the vectors $\{v_1, \dots, v_k\}$ defined by (1) can be chosen as the orthonormal eigenvectors associated with the k largest eigenvalues of the empirical covariance matrix Σ_n .

(5, 6, 7). (v_1, \dots, v_p) perech perepn' amaria $\approx \lambda_1, \dots, \lambda_p$. ut $\mathbf{\Sigma}_n$.

(v_1, \dots, v_d) : BON d'seck perepn' de Σ_n .

$$\Rightarrow \mathbf{v}^T \Sigma_n \mathbf{v} = \mathbf{v}^T \left[\sum_{i=1}^d \beta_i \underbrace{\sum_{j=1}^n v_j}_{\lambda_i v_i} \right] = \left(\sum_{h=1}^d \beta_h v_h \right)^T \left(\sum_{i=1}^d \beta_i \lambda_i v_i \right).$$

$$= \sum_{i=1}^d \beta_i^2 \lambda_i \leq \lambda_1 \underbrace{\sum_{i=1}^d \beta_i^2}_{\|\mathbf{v}\|^2=1} = \lambda_1.$$

$$\text{or: } \lambda_1^T \Sigma_n \mathbf{v}_1 = \lambda_1 \mathbf{v}_1^T \mathbf{v}_1 = \lambda_1.$$

Dmri: $v_1 \in \operatorname{argmax}_{\substack{\mathbf{v} \in \mathbb{R}^d \\ \|\mathbf{v}\|=1}} \mathbf{v}^T \Sigma_n \mathbf{v}$

$$\mathbf{v} = \sum_{i=1}^d \beta_i \mathbf{v}_i \quad \text{et } \mathbf{v} \in \operatorname{Span}(v_1, \dots, v_{d-1})^\perp$$

$$\begin{cases} \mathbf{v}^T \Sigma_n \mathbf{v} = \sum_{i=1}^d \beta_i^2 \lambda_i \leq \lambda_d \left(\sum_{i=1}^d \beta_i^2 \right) = \lambda_d. \\ \mathbf{v}_d^T \Sigma_n \mathbf{v}_d = \lambda_d \mathbf{v}_d^T \mathbf{v}_d = \lambda_d. \end{cases}$$

Dmri: $v_\ell \in \operatorname{argmax}_{\substack{\mathbf{v} \in \mathbb{R}^d \\ \|\mathbf{v}\|=1}} \mathbf{v}^T \Sigma_n \mathbf{v}$

$$\mathbf{v} \in \operatorname{Span}(v_1, \dots, v_{\ell-1})^\perp$$

8. The orthonormal eigenvectors associated with the eigenvalues of Σ_n allow to define the principal components as follows. Then, as $V_d = \text{span}\{\vartheta_1, \dots, \vartheta_d\}$, for all $1 \leq i \leq n$,

$$\pi_{V_d}(X_i) = \sum_{k=1}^d \langle X_i, \vartheta_k \rangle \vartheta_k = \sum_{k=1}^d (X_i^\top \vartheta_k) \vartheta_k = \sum_{k=1}^d c_k(i) \vartheta_k,$$

where for all $1 \leq k \leq d$, the k -th principal component is defined as $c_k = \mathbf{X} \vartheta_k$. Prove that (c_1, \dots, c_d) are orthogonal vectors.

$$\pi_{V_d}(X_i) = P_{V_p}^\perp(X_i) = \sum_{k=1}^d \underbrace{\langle X_i, v_k \rangle}_{\underbrace{X_i^\top v_k}_{c_k(i)}} v_k.$$

$$c_k(i) = X_i^\top v_k.$$

$$\forall k \in \{1, \dots, d\} \quad \in \Sigma_m.$$

$$c_k = \begin{pmatrix} c_k(1) \\ \vdots \\ c_k(n) \end{pmatrix} = \underbrace{\begin{pmatrix} X_1^\top \\ \vdots \\ X_n^\top \end{pmatrix}}_X v_k.$$

$$\langle c_h, c_l \rangle = 0 = v_h^\top \underbrace{X^\top X}_{\propto \lambda_l} v_l = \propto \lambda_l \underbrace{v_h^\top v_l}_{=0 \text{ if } h \neq l}.$$

EXERCISE 3 (KERNEL PRINCIPAL COMPONENT ANALYSIS) Let $(X_i)_{1 \leq i \leq n}$ be n observations in a general space \mathcal{X} , $k : \mathcal{X} \times \mathcal{X} \rightarrow \mathbb{R}$ a positive kernel and $\mathbf{K} = (k(X_i, X_j))_{1 \leq i, j \leq n}$. \mathcal{W} denotes the Reproducing Kernel Hilbert Space associated with k and for all $x \in \mathcal{X}$, $\phi(x)$ denotes the function $\phi(x) : y \mapsto k(x, y)$. The aim is now to perform a PCA on $(\phi(X_1), \dots, \phi(X_n))$. It is assumed that $\sum_{i=1}^n \phi(X_i) = 0$.

1. Prove that

$$f_1 = \underset{f \in \mathcal{W}; \|f\|_{\mathcal{W}}=1}{\operatorname{argmax}} \sum_{i=1}^n \langle \phi(X_i), f \rangle_{\mathcal{W}}^2$$

may be written

$$f_1 = \sum_{i=1}^n \alpha_1(i) \phi(X_i), \quad \text{where } \alpha_1 = \underset{\alpha \in \mathbb{R}^n; \alpha^\top \mathbf{K} \alpha = 1}{\operatorname{argmax}} \alpha^\top \mathbf{K}^2 \alpha.$$

2. Prove that $\alpha_1 = \lambda_1^{-1/2} b_1$ where b_1 is the unit eigenvector associated with the largest eigenvalue λ_1 of \mathbf{K} . What about (f_2, \dots, f_p) defined iteratively as in (1)?

3. Write $H_p = \text{span}\{f_1, \dots, f_p\}$. Prove that, for all $1 \leq i \leq n$,

$$\pi_{H_p}(\phi(X_i)) = \sum_{j=1}^p \lambda_j \alpha_j(i) f_j .$$